Personalized screening intervals for biomarkers using joint models for longitudinal and survival data

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- Nowadays growing interest in tailoring medical decision making to individual patients
 - Personalized Medicine
 - > Shared Decision Making
- This is of high relevance in various diseases
 - ▷ cancer research, cardiovascular diseases, HIV research, ...

Physicians are interested in accurate prognostic tools that will inform them about the future prospect of a patient in order to adjust medical care



- Aortic Valve study: Patients who received a human tissue valve in the aortic position
 - b data collected by Erasmus MC (from 1987 to 2008);
 77 received sub-coronary implantation; 209 received root replacement

• Outcomes of interest:

- \triangleright death and re-operation \rightarrow composite event
- ▷ aortic gradient



• General Questions:

Can we utilize available aortic gradient measurements to predict survival/re-operation?

> When to plan the next echo for a patient?



• Goals of this talk:

- \triangleright introduce joint models
- ▷ dynamic predictions
- \triangleright optimal timing of next visit



- To answer these questions we need to postulate a model that relates
 - \triangleright the aortic gradient with
 - ▷ the time to death or re-operation
- Some notation
 - $\triangleright T_i^*$: True time-to-death for patient *i*
 - $\triangleright T_i$: Observed time-to-death for patient i
 - $\triangleright \delta_i$: Event indicator, i.e., equals 1 for true events
 - $\triangleright y_i$: Longitudinal aortic gradient measurements

2.1 Joint Modeling Framework (cont'd)







- We start with a standard joint model
 - ▷ Survival Part: Relative risk model

$$h_i(t \mid \mathcal{M}_i(t)) = h_0(t) \exp\{\gamma^\top w_i + \alpha m_i(t)\},\$$

where

* $m_i(t)$ = the *true* & *unobserved* value of aortic gradient at time t* $\mathcal{M}_i(t) = \{m_i(s), 0 \le s < t\}$ * α quantifies the effect of aortic gradient on the risk for death/re-operation * w_i baseline covariates



 $\triangleright \text{ Longitudinal Part: } \text{Reconstruct } \mathcal{M}_i(t) = \{m_i(s), 0 \leq s < t\} \text{ using } y_i(t) \text{ and a mixed effects model (we focus on continuous markers)}$

$$y_i(t) = m_i(t) + \varepsilon_i(t)$$

$$= x_i^{\top}(t)\beta + z_i^{\top}(t)b_i + \varepsilon_i(t), \qquad \varepsilon_i(t) \sim \mathcal{N}(0, \sigma^2),$$

where

* $x_i(t)$ and β : Fixed-effects part * $z_i(t)$ and b_i : Random-effects part, $b_i \sim \mathcal{N}(0, D)$



- \bullet The two processes are associated \Rightarrow define a model for their joint distribution
- Joint Models for such joint distributions are of the following form (Tsiatis & Davidian, *Stat. Sinica*, 2004; Rizopoulos, CRC Press, 2012)

$$p(y_i, T_i, \delta_i) = \int p(y_i \mid b_i) \{ h(T_i \mid b_i)^{\delta_i} S(T_i \mid b_i) \} p(b_i) db_i$$

where

 $\triangleright b_i$ a vector of random effects that explains the interdependencies $\triangleright p(\cdot)$ density function; $S(\cdot)$ survival function



- Joint models can be estimated with either Maximum Likelihood or Bayesian approaches (i.e., MCMC)
- Here we follow the Bayesian approach because it facilitates computations for our later developments...



- We are interested in predicting survival probabilities for a new patient j that has provided a set of aortic gradient measurements up to a specific time point t
- Example: We consider Patients 20 and 81 from the Aortic Valve dataset















- What do we know for these patients?
 - ▷ a series of aortic gradient measurements
 - ▷ patient are event-free up to the last measurement
- **Dynamic Prediction** survival probabilities are dynamically updated as additional longitudinal information is recorded



- \bullet <u>Available info:</u> A new subject j with longitudinal measurements up to t
 - $\triangleright T_j^* > t$ $\triangleright \mathcal{Y}_j(t) = \{y_j(t_{jl}); 0 \le t_{jl} \le t, l = 1, \dots, n_j\}$
 - $\triangleright \mathcal{D}_n$ sample on which the joint model was fitted





• Based on the fitted model we can estimate the conditional survival probabilities

$$\pi_j(u \mid t) = \mathsf{Pr}\big\{T_j^* \ge u \mid T_j^* > t, \mathcal{Y}_j(t), \mathcal{D}_n\big\}, \quad u > t$$

- For more details check:
 - Proust-Lima and Taylor (2009, Biostatistics), Rizopoulos (2011, Biometrics), Taylor et al. (2013, Biometrics)



- Example: We fit a joint model to the Aortic Valve data
- Longitudinal submodel
 - \triangleright fixed effects: natural cubic splines of time (d.f.= 3), operation type, and their interaction
 - \triangleright random effects: Intercept, & natural cubic splines of time (d.f.= 3)
- Survival submodel
 - \triangleright type of operation, age, sex + underlying aortic gradient level
 - ▷ log baseline hazard approximated using B-splines































• Question 2:

> When the patient should come for the next visit?



This is a difficult question!

- Many parameters that affect it
 - ▷ which model to use?
 - ▷ what criterion to use?
 - ▷ change in treatment?
 - $\triangleright \dots$

We will work under the following setting \Rightarrow















- Let $y_i(u)$ denote the future longitudinal measurement u > t
- We would like to select the optimal u such that:
 - \triangleright patient still event-free up to u
 - \triangleright maximize the information by measuring $y_j(u)$ at u



• Utility function

$$U(u \mid t) = E\left\{\lambda_{1} \underbrace{\log \frac{p\left(T_{j}^{*} \mid T_{j}^{*} > u, \left\{\mathcal{Y}_{j}(t), \boldsymbol{y}_{j}(\boldsymbol{u})\right\}, \mathcal{D}_{n}\right)}{p\left\{T_{j}^{*} \mid T_{j}^{*} > u, \mathcal{Y}_{j}(t), \mathcal{D}_{n}\right\}}}_{First term} + \lambda_{2} \underbrace{I(T_{j}^{*} > \boldsymbol{u})}_{Second term}\right\}$$

expectation wrt joint predictive distribution $[T_j^*, y_j(u) \mid T_j^* > t, \mathcal{Y}_j(t), \mathcal{D}_n]$

- \triangleright First term: expected Kullback-Leibler divergence of posterior predictive distributions with and without $y_i(u)$
- \triangleright Second term: 'cost' of waiting up to $u \Rightarrow$ increase the risk



- \bullet Nonnegative constants λ_1 and λ_2 weigh the cost of waiting as opposed to the information gain
 - \triangleright elicitation in practice difficult \Rightarrow trading information units with probabilities
- How to get around it?

Equivalence between compound and constrained optimal designs



- It can be shown that
 - \triangleright for any λ_1 and λ_2 ,

 \triangleright there exists a constant $\pmb{\kappa} \in [0,1]$ for which

$$\underset{u}{\operatorname{argmax}} \bigcup(u \mid t) \iff \underset{u}{\operatorname{argmax}} E\left\{ \log \frac{p(T_j^* \mid T_j^* > u, \{\mathcal{Y}_j(t), y_j(u)\}, \mathcal{D}_n)}{p\{T_j^* \mid T_j^* > u, \mathcal{Y}_j(t), \mathcal{D}_n\}} \right\}$$

subject to the constraint $\pi_j(u \mid t) \geq \kappa$



- Elicitation of κ is relatively easier
 - ▷ Chosen by the physician
 - Determined using ROC analysis
- Estimation is achieved using a Monte Carlo scheme
 - ▷ more details in Rizopoulos et al. (2015)



Example: We illustrate how for Patient 81 we have seen before
 The threshold for the constraint is set to

$$\pi_j(u \mid t) \ge \kappa = 0.8$$

> After each visit we calculate the optimal timing for the next one using

$$\underset{u}{\operatorname{argmax}} \operatorname{\mathsf{EKL}}(u \mid t) \quad \text{ where } \ u \in (t, t^{up}]$$

and

$$t^{up} = \min\{5, u : \pi_j(u \mid t) = 0.8\}$$

























5. Software



- Software: R package **JMbayes** freely available via http://cran.r-project.org/package=JMbayes
 - \triangleright it can fit a variety of joint models + many other features
 - > relevant to this talk: cvDCL() and dynInfo()

GUI interface for dynamic predictions using package shiny

Thank you for your attention!