# Using Joint Models to Estimate Causal Effects for Salvage Therapy after Prostatectomy

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# Aims, Models & Estimands

## 1 Background & Aim



- Setting Patients treated with surgery after diagnosis of Prostate Cancer (PCa)
  - > remain at risk of metastasis
- Follow-up
  - > PSA levels at frequent intervals

  - > ST androgen deprivation therapy, radiation therapy, chemotherapy, and combinations



- Important questions regarding Salvage Therapy

  - ▶ when to start?
  - ▷ does it work?



Quantify the amount by which Salvage Therapy reduces the risk of metastasis



## University of Michigan Prostatectomy Data

- → 3634 PCa patients followed-up in 1996–2013
  - \* aged 40 to 84 years with clinically localized cT1 to cT3 disease
  - \* received radical prostatectomy
- baseline variables: PSA, Gleason, T-stage, age, race, gland volume, perineural invasion, planned adjuvant therapy



## Challenges

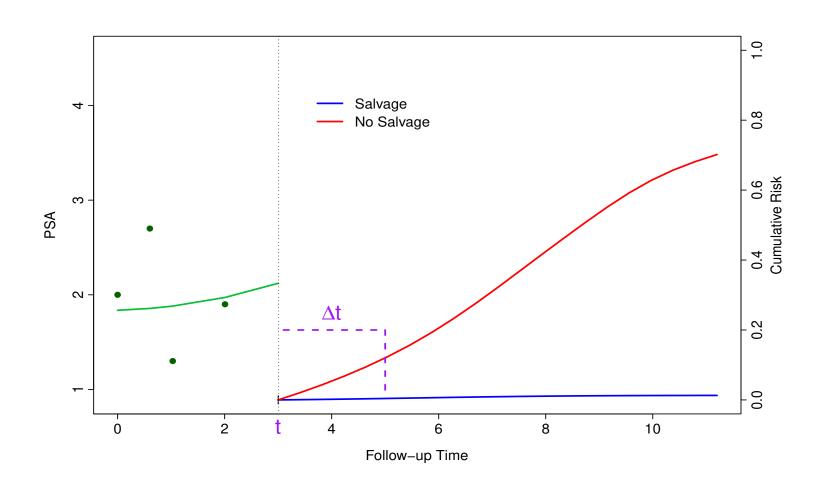
- ▷ Observational Data no RCT
  - \* selection bias
  - \* ascertainment bias
- ▷ Time-Varying Salvage Therapy
  - \* depends on previous PSA
  - \* PSA time-dependent confounder
  - \* endogeneity

#### 2 Causal ST Effects



- Standard assumptions for Causal Inference
  - Consistency: Observed outcomes equal the counterfactual outcomes for the actually assigned treatment
  - ▷ Sequential Exchangeability: The counterfactual outcomes are independent of the assigned treatment conditionally on the history of PSA measurements and baseline covariates







#### Which is the target group?

#### Notation

 $\triangleright T_m$ : time to metastasis

 $\triangleright T_d$ : time to death

 $\triangleright \mathcal{H}^*(t)$ : a version of the PSA history up to t

 $hd T_m^{(a)}$  and  $T_d^{(a)}$  counterfactual outcomes

\* a=1, ST given at t

\* a=0, ST was not given in  $[t,t+\Delta t]$ 



## Marginal Salvage Therapy Effect

b we average over all PSA histories

$$ST^{M}(t + \Delta t, t) = \Pr\{T_{m}^{(1)} \le t + \Delta t \mid T_{m} > t, T_{d} > t\} - \Pr\{T_{m}^{(0)} \le t + \Delta t \mid T_{m} > t, T_{d} > t\}$$

#### • Notes:

 $\triangleright$  of lesser relevance to the urologists because they decide who gets ST based on PSA  $\Rightarrow$  more bias

 $\triangleright$  averages over a big group of patients  $\Rightarrow$  smaller variance



## Conditional Salvage Therapy Effect

 $\triangleright$  we condition on the PSA history of a specific patient, i.e.,  $\mathcal{H}^*(t) = \mathcal{H}_i(t)$ 

$$ST^{C}(t + \Delta t, t) = \Pr\{T_{m}^{(1)} \le t + \Delta t \mid T_{m} > t, T_{d} > t, \mathcal{H}_{i}(t)\}$$
$$-\Pr\{T_{m}^{(0)} \le t + \Delta t \mid T_{m} > t, T_{d} > t, \mathcal{H}_{i}(t)\}$$

#### • Notes:

 $\triangleright$  much more relevant to the urologists  $\Rightarrow$  **less bias** 

 $\triangleright$  averages over a narrow group of patients  $\Rightarrow$  larger variance



# Marginal-Conditional Salvage Therapy Effect

 $\triangleright$  consider ST for patients who had PSA levels above the threshold value c at their last visit, i.e.,  $\mathcal{H}^*(t)=\{Y(t):Y(t)>c\}$ 

$$ST^{MC}(t + \Delta t, t) = \Pr\{T_m^{(1)} \le t + \Delta t \mid T_m > t, T_d > t, \mathcal{H}^*(t)\}$$
$$-\Pr\{T_m^{(0)} \le t + \Delta t \mid T_m > t, T_d > t, \mathcal{H}^*(t)\}$$

#### • Notes:

- ▷ relevant to the urologists ⇒ compromised bias
- ▷ averages over a bigger group of patients ⇒ compromised variance

## 3 Structural Models



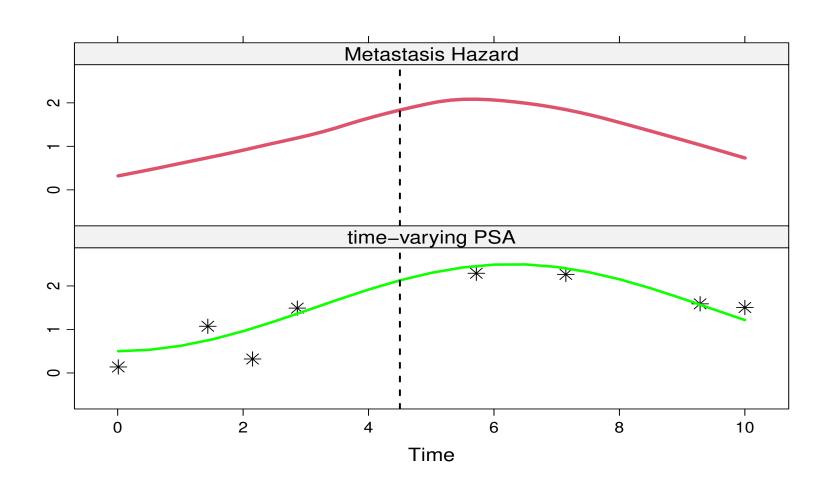
Standard Cox models not appropriate



Joint Models for Longitudinal and Time-to-Event Data

# 3 Structural Models (cont'd)





## 3 Structural Models (cont'd)



Joint models completely specify the joint distribution of PSA, time-to-metastasis & time-to-death

- Under sequential ignorability,
  - > they provide valid marginal distributions
  - > without requiring to model the treatment assignment mechanism

## 4 PSA Sub-Model



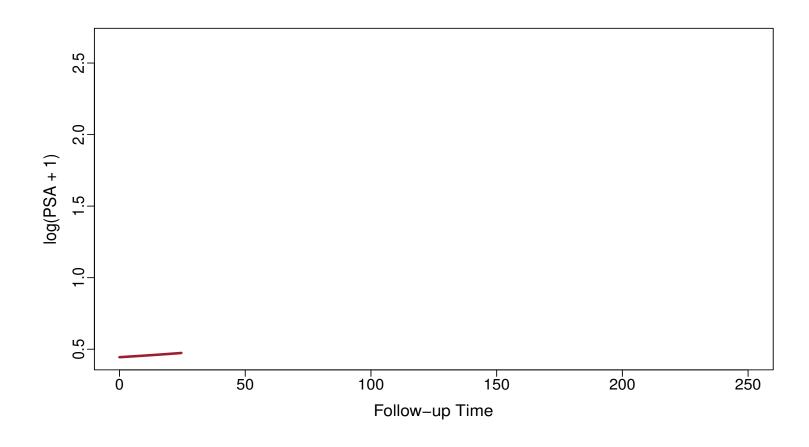
- As PSA increases, patients may receive ST
- ullet We let  $S_i$  denote the time a patient initiated ST
  - $\triangleright$  for patients who did not initiate ST,  $S_i = \infty$
- After ST, PSA levels are expected to drop
  - but may rise again before metastasis



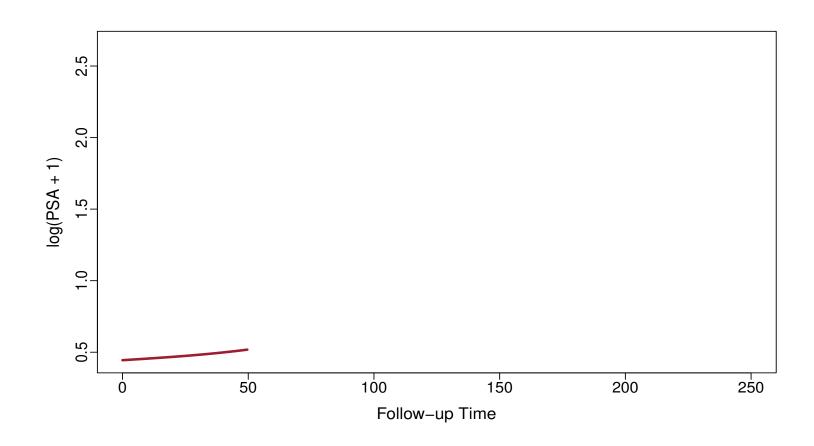
$$\log\{\mathsf{PSA}_i(t)+1\} = \begin{cases} \eta_i(t) + \varepsilon_i(t) = \boldsymbol{x}_i(t)\boldsymbol{\beta} + \boldsymbol{z}_i(t)\boldsymbol{b}_i + \varepsilon_i(t), & t < S_i \\ \\ \tilde{\eta}_i(t) + \varepsilon_i(t) = \\ \\ \eta_i(t) + \left\{\tilde{\boldsymbol{x}}_i(\tilde{t})\tilde{\boldsymbol{\beta}} + \tilde{\boldsymbol{z}}_i(t)\tilde{\boldsymbol{b}}_i\right\} + \varepsilon_i(t), & t \geq S_i, \end{cases}$$

$$oldsymbol{u}_i = (oldsymbol{b}_i, \widetilde{oldsymbol{b}}_i) \sim \mathcal{N}(oldsymbol{0}, oldsymbol{\Omega})$$

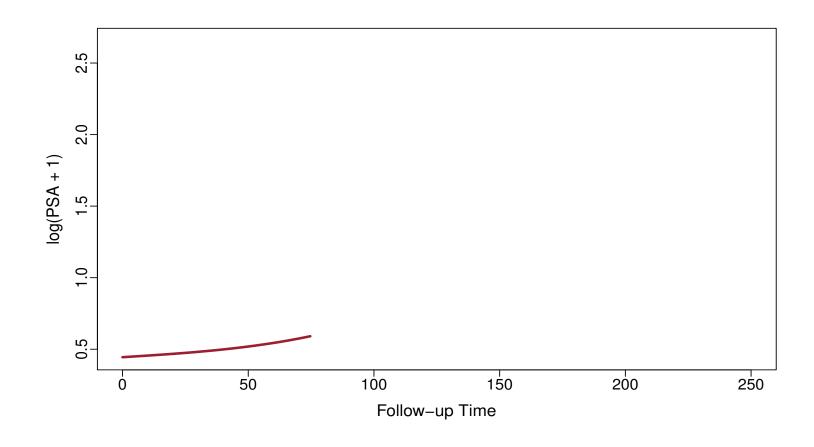




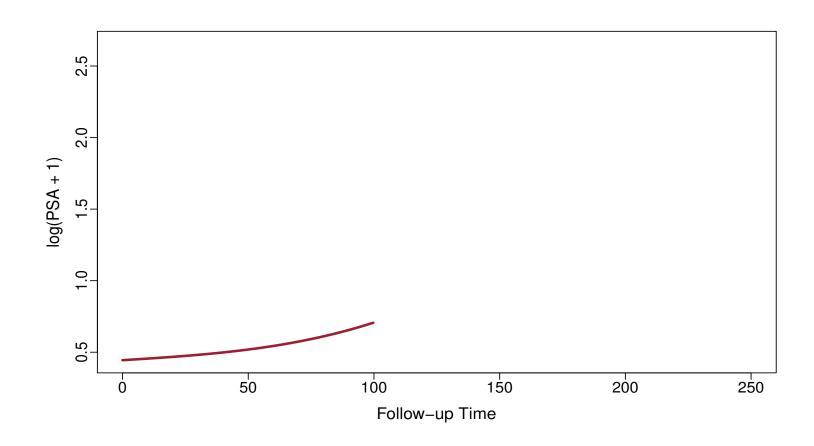




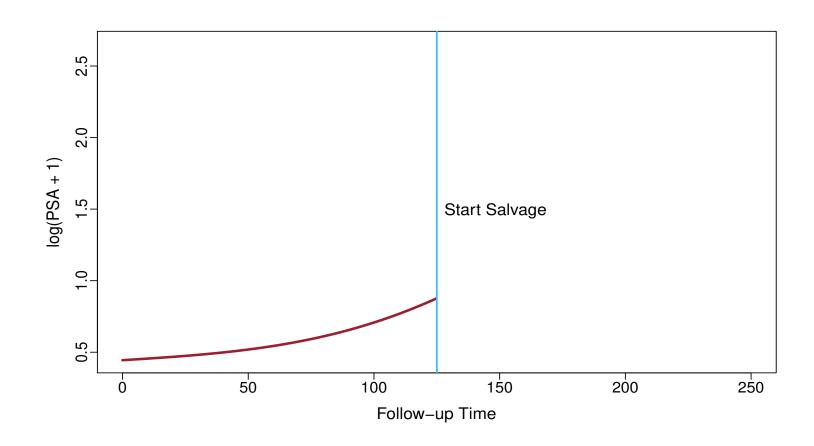




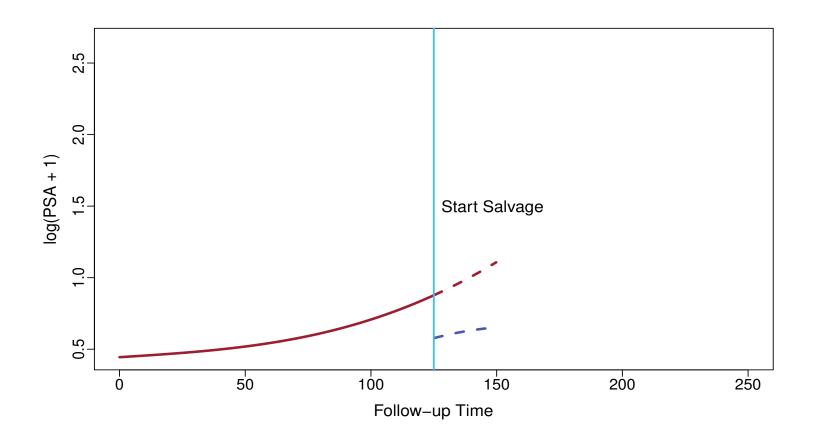




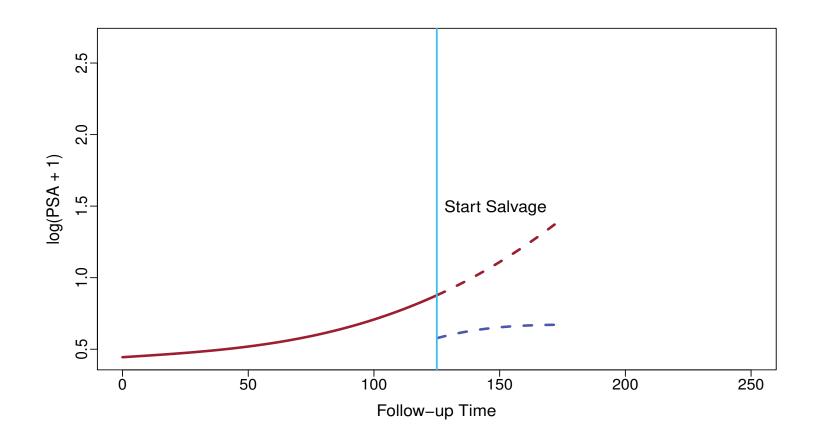




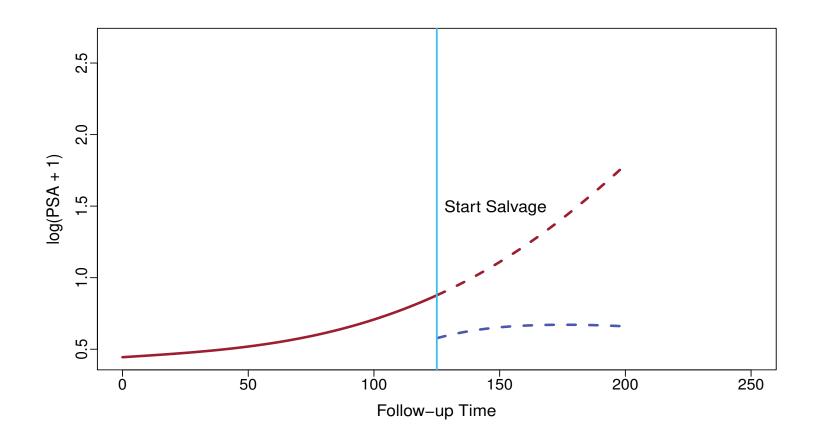




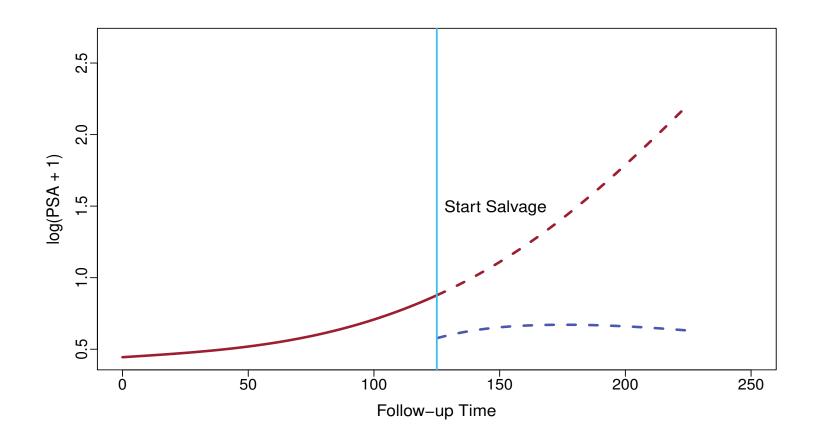




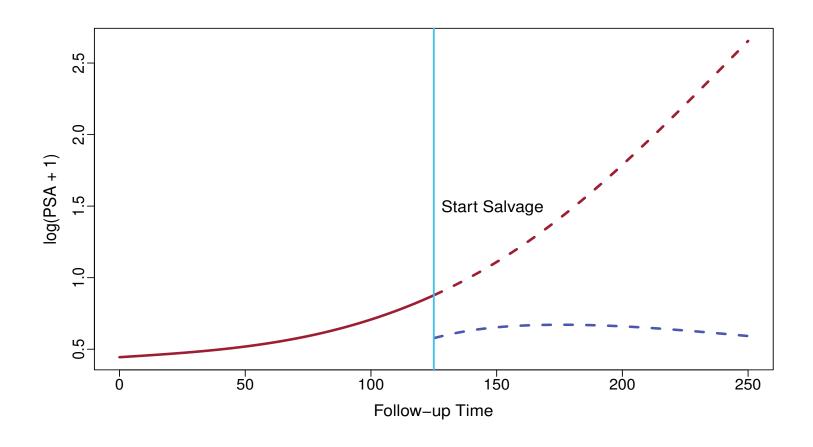












### 5 Metastasis and Death Sub-Models



- Metastasis and Death treated as *Competing Risks*
- Separate hazard models for metastasis and death

  - ▷ baseline covariates

## 5 Metastasis and Death Sub-Models (cont'd)



Metastasis Sub-Model linked to baseline covariates, Salvage and PSA

$$h_i^m(t) = \begin{cases} h_0^m(t) \exp\left(\boldsymbol{\psi}_m^{\top} \boldsymbol{w}_i + \boldsymbol{\alpha}_m^{\top} f\{\eta_i(t)\}\right), & t < S_i \\ h_0^m(t) \exp\left(\boldsymbol{\psi}_m^{\top} \boldsymbol{w}_i + \gamma_m(t - S_i) + \boldsymbol{\xi}_m^{\top} g\{\tilde{\eta}_i(t)\}\right), & t \ge S_i \end{cases}$$

## 5 Metastasis and Death (cont'd)



• Death Sub-Model linked to baseline covariates, Salvage but not PSA

$$h_i^d(t) = \begin{cases} h_0^d(t) \exp(\boldsymbol{\psi}_d^{\top} \boldsymbol{w}_i), & t < S_i \\ h_0^d(t) \exp(\boldsymbol{\psi}_d^{\top} \boldsymbol{w}_i + \gamma_d), & t \ge S_i \end{cases}$$

### 6 Causal Effect Estimation



• From the joint model, we can obtain the conditional causal effect

$$\Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \mathcal{H}_i(t), \mathcal{X}_i\} =$$

$$\int \int \Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \boldsymbol{u}_i, \mathcal{X}_i, \boldsymbol{\theta}\}$$

$$\times p\{\boldsymbol{u}_i \mid T_{mi} > t, T_{di} > t, \mathcal{H}_i(t), \mathcal{X}_i, \boldsymbol{\theta}\} \ p(\boldsymbol{\theta} \mid \mathcal{D}) \ d\boldsymbol{u}_i d\boldsymbol{\theta}$$

## 6 Causal Effect Estimation (cont'd)



- ullet Monte Carlo scheme to estimate  ${\sf ST}_i^C(t+\Delta t,t)$ 
  - riangle sample  $reve{m{ heta}}^{(l)}$  from the posterior of the parameters  $[m{ heta} \mid \mathcal{D}]$
  - ightharpoonup sample  $m{ar{u}}_i^{(l)}$  from the posterior of the random effects  $[m{u}_i \mid T_{mi} > t, T_{di} > t, \mathcal{H}_i(t), \mathcal{X}_i, m{ar{ heta}}^{(l)}]$

$$ho$$
 calculate  $\pi_i^{(l)}(t + \Delta t \mid t, a) = \Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \boldsymbol{\check{u}}_i^{(l)}, \boldsymbol{\mathcal{X}}_i, \boldsymbol{\check{\theta}}^{(l)}\}$ 

ullet We repeat L times and get

$$\widehat{\mathsf{ST}}_i^C(t + \Delta t, t) = \frac{1}{L} \sum_{l=1}^L \pi_i^{(l)}(t + \Delta t \mid t, a = 1) - \pi_i^{(l)}(t + \Delta t \mid t, a = 0)$$

## 6 Causal Effect Estimation (cont'd)

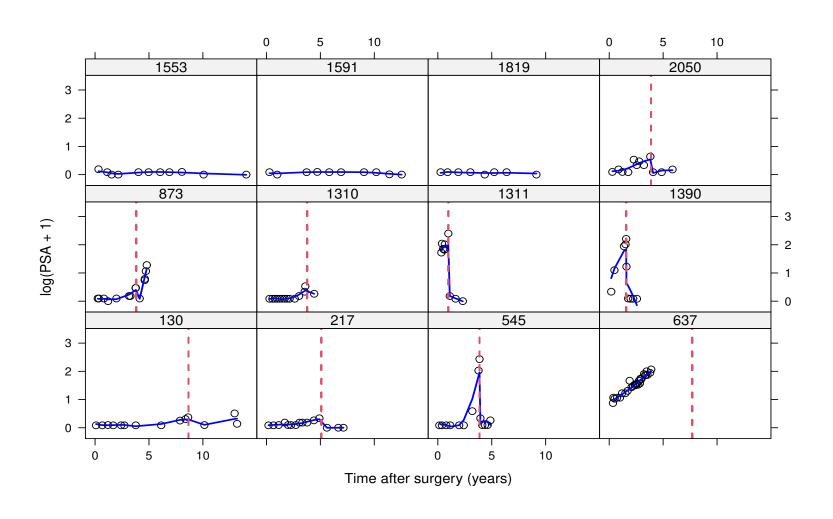


- ullet Estimation of  $\mathrm{ST}^M(t+\Delta t,t)$  and  $\mathrm{ST}^{MC}(t+\Delta t,t)$  proceeds by averaging the conditional effects over the respective groups of patients
- ullet For example, for  $\mathrm{ST}^M(t+\Delta t,t)$ 
  - $\triangleright \mathcal{R}(t)$  the subset of patients at risk at time t
  - ho for each patient in  $\mathcal{R}(t)$ , we calculate  $\widehat{\mathsf{ST}}_i^C(t+\Delta t,t)$

$$\widehat{\mathsf{ST}}^M(t+\Delta t,t) = n_r^{-1} \sum_{i:i \in R(t)} \widehat{\mathsf{ST}}_i^C(t+\Delta t,t),$$

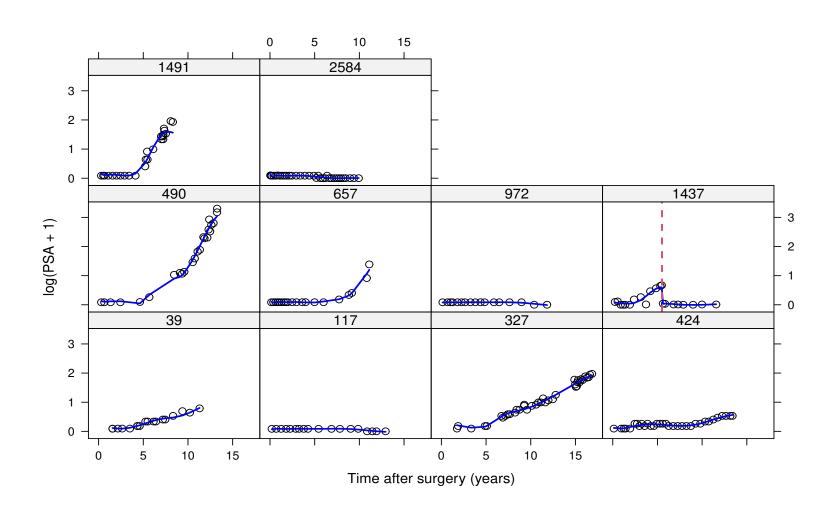
## 7 Results





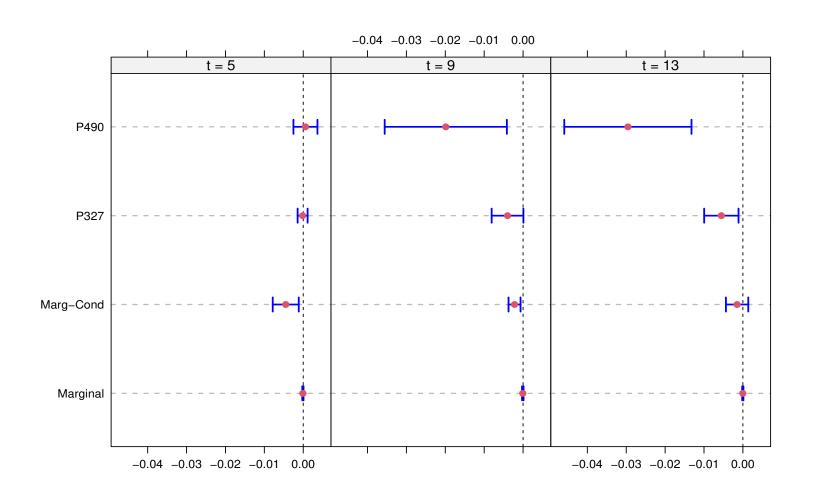
## 7 Results (cont'd)





# 7 Results (cont'd)





## 7 Software (cont'd)



- Implementation available in JMbayes2
  - > predict() cumulative incidence risks
- Shiny app...

## Thank for your attention!

https://www.drizopoulos.com/